

FALL/2018/MA526: MID-TERM I SOLUTION

Instructor: Guangqu Zheng¹

Wednesday, September 26: 2:00-2:45pm

Maximal points: 20.

Attention: You need to provide enough supporting arguments for your response.

Q1 (4pt) Let x_1, \dots, x_5 be a sample data.

- (1) Recall the definition of its sample average. (1pt)
- (2) Recall the definition of its sample median. (1pt)
- (3) Is sample average always bigger than sample median? (2pt) Argue with examples.

Answer: (1) In this question, there are only five numbers, so its sample average is given by

$$(x_1 + x_2 + x_3 + x_4 + x_5)/5$$

(2) We first relabel the numbers as y_1, y_2, y_3, y_4, y_5 such that $y_1 \leq y_2 \leq y_3 \leq y_4 \leq y_5$. If so, the median of x_1, \dots, x_5 is y_3 .

(3) No. The sample average may be bigger than sample median: For example $x_1 = x_2 = x_3 = x_4 = 0$, $x_5 = 5$, then the sample average is 1 while the median is zero.

The sample average may be equal to its sample median: For example $x_1 = x_2 = x_3 = x_4 = x_5 = 5$, then the sample average the median are both equal to 5.

The sample average may be smaller than its sample median: For example $x_1 = x_2 = x_3 = x_4 = 5$ and $x_5 = 0$, then the sample average is 4 & the median is 5.

To conclude, the sample average is not always bigger than the sample median.

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Q2 (4pt) Suppose you roll two fair 6-sided dice. What is the probability that the sum is at least 8 given that you rolled at least one 5?

Answer: We are looking for the conditional probability

$$\text{Prob}(\text{the sum is at least 8} \mid \text{you rolled at least one 5})$$

which is equal to, by definition,

$$\frac{\text{Prob}(\text{the sum is at least 8 and you rolled at least one 5})}{\text{Prob}(\text{you rolled at least one 5})}$$

Now back to the usual argument: when we roll two fair dice, we get 36 possible outcomes, each of which has equal probability $1/36$:

$$\Omega := \{(1, 1), (1, 2), \dots, (1, 6), (2, 1), (2, 2), \dots, (2, 6), \dots, (6, 6)\}$$

The event {the sum is at least 8 **and** you rolled at least one 5} contains the following outcomes:

$$\{(3, 5), (4, 5), (5, 5), (6, 5), (5, 3), (5, 4), (5, 6)\}$$

this event has probability $7/36$.

The event {you rolled at least one 5} contains the following outcomes:

$$\{(1, 5), (2, 5), (3, 5), (4, 5), (5, 5), (6, 5), (5, 1), (5, 2), (5, 3), (5, 4), (5, 6)\}$$

this event has probability $11/36$.

So the required condition probability is $7/11$.

Q3-1 (3pt) Determine the value C so that the following function can serve as a probability density function of the random variable X :

$$f(x) = Cx^2 \text{ for } 0 \leq x \leq 1; f(x) = C|x| \text{ for } -2 \leq x \leq -1; f(x) = 0 \text{ elsewhere.}$$

Answer: First it is clear that we require $C > 0$. And we also need

$$1 = \int_{-\infty}^{\infty} f(x) dx = \int_0^1 Cx^2 dx + \int_{-2}^{-1} C|x| dx$$

Note that

$$\int_0^1 Cx^2 dx = C \int_0^1 x^2 dx = C \left(\frac{x^3}{3} \right) \Big|_{x=0}^{x=1} = C/3$$

while (as $|x| = -x$ for $x < 0$)

$$\int_{-2}^{-1} C|x| dx = -C \int_{-2}^{-1} x dx = -C \left(\frac{x^2}{2} \right) \Big|_{x=-2}^{x=-1} = \frac{3}{2}C$$

Hence $1 = \frac{C}{3} + \frac{3}{2}C = \frac{11}{6}C$ implying $C = 6/11$.

Q3-(2)(4pt) Two fair dice are rolled. Consider the events $A = \{ \text{sum of two dice equals 3} \}$,
 $B = \{ \text{sum of two dice equals 7} \}$ and $C = \{ \text{at least one of the dice shows a 1} \}$.

- (a) What is $\mathbb{P}(A|C)$? (1pt)
 (b) What is $\mathbb{P}(B|C)$? (1pt)
 (c) Are event A and event C independent? What about B and C ? (2pt)

Answer :

(a) $\mathbb{P}(A|C) = \mathbb{P}(A \cap C)/\mathbb{P}(C)$ and the event $A \cap C$ means sum of two dice is three **and** at least one of the dice shows 1, that it is equivalent to say that $(1, 2)$ or $(2,1)$ are rolled. So $\mathbb{P}(A \cap C) = 2/36 = 1/18$. (Note that in this question, we have $A \subset C$)

And the event C means that one of $(1, 1), (1, 2), \dots, (1, 6), (2, 1), (3, 1), (4, 1), (5, 1), (6, 1)$ is rolled, so $\mathbb{P}(C) = 11/36$.

Thus, $\mathbb{P}(A|C) = \mathbb{P}(A \cap C)/\mathbb{P}(C) = 2/11$.

(b) $\mathbb{P}(B|C) = \mathbb{P}(B \cap C)/\mathbb{P}(C)$ and the event $B \cap C$ means sum of two dice is 7 **and** at least one of the dice shows 1, that it is equivalent to say that $(1, 6)$ or $(6,1)$ are rolled. So $\mathbb{P}(B \cap C) = 2/36 = 1/18$.

And we know from part (a) that $\mathbb{P}(C) = 11/36$. Thus, $\mathbb{P}(B|C) = \mathbb{P}(B \cap C)/\mathbb{P}(C) = 2/11$.

(c) The event A and event C are **not independent**, the event B and event C are **not independent** either: We know from part (a) that $\mathbb{P}(A|C) = \mathbb{P}(A)/\mathbb{P}(C) \neq \mathbb{P}(A)$, as $A \subset C$; and we know from part (b) that $\mathbb{P}(B|C) = 2/11$ while $\mathbb{P}(B) = 6/36$.

The result $\mathbb{P}(B) = 6/36$ comes from the fact that the event B means that one of $(1, 6), (6, 1), (2, 5), (5, 2), (3, 4), (4, 3)$ is rolled.

Q4: Next page

Bonus question (3 pt) Let X be a random variable with finite second moment and write \mathbb{E} for the expectation. Then what is the minimal value for the function $\phi(x) := \mathbb{E}([X - x]^2)$?

Answer :

We know $(y - x)^2 = y^2 - 2xy + x^2$, so $[X - x]^2 = X^2 - 2xX + x^2$. Therefore,

$$\phi(x) := \mathbb{E}([X - x]^2) = \mathbb{E}(X^2 - 2xX + x^2) = \mathbb{E}[X^2] - 2\mathbb{E}[X]x + x^2$$

This is a function in the variable x , while $\mathbb{E}[X^2]$ and $2\mathbb{E}[X]$ are just numbers. It is enough to find x_0 such that $\phi'(x_0) = 0$, as we know $\phi(x)$ has a unique minimum. Let us compute the derivative of $\phi(x)$ now:

$$\frac{d}{dx}\phi(x) = 2x - 2\mathbb{E}[X]$$

so that we should take $x_0 = \mathbb{E}[X]$ and $\phi(x_0) = \mathbb{E}([X - \mathbb{E}(X)]^2) = \text{Var}(X)$ is the minimal value of the function ϕ .

Q4 (5pt) Let Ω be a sample space equipped with a probability \mathbb{P} . Let A, B be two events on Ω such that A, B are independent and $\mathbb{P}(A) = \mathbb{P}(B) = 1/2$. Now we define the random variable $X(\omega) = \mathbb{I}_A(\omega) + \mathbb{I}_B(\omega)$, where the indicator function \mathbb{I}_A is defined by

$$\mathbb{I}_A(\omega) = 1 \quad \text{if } \omega \in A \quad \text{and} \quad \mathbb{I}_A(\omega) = 0 \quad \text{if } \omega \text{ is not in } A.$$

- (1) Find $\mathbb{P}(A \cap B)$ (1pt) and find $\mathbb{P}(A^c \cap B^c)$ (1pt)
- (2) Write down the probability mass function of X . (1pt)
- (3) Write down the cumulative distributional function of X . (1pt)
- (4) Compute the expectation of X . (1pt)

Answer:

(1) Since A, B are independent, we have $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B) = 1/4$. By de-Morgan's law $(A \cup B)^c = A^c \cap B^c$, then

$$\mathbb{P}(A^c \cap B^c) = 1 - \mathbb{P}(A \cup B) = 1 - [\mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)]$$

where the last equality follows from the additive rules. So we have $\mathbb{P}(A^c \cap B^c) = 1 - [1/2 + 1/2 - 1/4] = 1/4$.

(2) It is clear to see the random variable X defined above is discrete with possible values in $\{0, 1, 2\}$. And we have

$$X(\omega) = 0 \quad \text{if and only if} \quad \mathbb{I}_A(\omega) = \mathbb{I}_B(\omega) = 0 \quad \text{if and only if} \quad \omega \in A^c \cap B^c$$

In other words, $\{X = 0\} = A^c \cap B^c$, therefore $\mathbb{P}(X = 0) = 1/4$. And similarly, we have

$$X(\omega) = 2 \quad \text{if and only if} \quad \mathbb{I}_A(\omega) = \mathbb{I}_B(\omega) = 1 \quad \text{if and only if} \quad \omega \in A \cap B$$

In other words, $\{X = 2\} = A \cap B$, therefore $\mathbb{P}(X = 2) = 1/4$. Hence

$$\mathbb{P}(X = 1) = 1 - \mathbb{P}(X = 0) - \mathbb{P}(X = 2) = 1/2$$

implying that the probability mass function f_X of the random variable X is given by

$$f_X(x) = \begin{cases} 1/4 & \text{if } x = 0 \\ 1/2 & \text{if } x = 1 \\ 1/4 & \text{if } x = 2 \\ 0 & \text{elsewhere.} \end{cases}$$

(3) The corresponding cumulative distribution function F_X is given by

$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1/4 & \text{if } 0 \leq x < 1 \\ 3/4 & \text{if } 1 \leq x < 2 \\ 1 & \text{if } x \geq 2. \end{cases}$$

(4) By definition,

$$\mathbb{E}[X] = \sum_x x f_X(x) = 0 \times \frac{1}{4} + 1 \times \frac{1}{2} + 2 \times \frac{1}{4} = 1.$$

Alternatively, we have $\mathbb{E}[X] = \mathbb{E}[\mathbb{I}_A + \mathbb{I}_B] = \mathbb{E}[\mathbb{I}_A] + \mathbb{E}[\mathbb{I}_B] = \mathbb{P}(A) + \mathbb{P}(B) = 1$.